

OUTLINE PROOF – MAXIMUM LIKELIHOOD METHOD

The likelihood is the probability of observing a particular dataset, therefore

$$\int L(x | \theta) dx = 1$$

differentiate with respect to θ

$$\int \frac{\partial L}{\partial \theta} dx = 0 = \int \frac{1}{L} \frac{\partial L}{\partial \theta} \cdot L dx = \int \frac{\partial \ln(L)}{\partial \theta} \cdot L dx$$

differentiate RH term with respect to θ again

$$\int \frac{\partial^2 \ln(L)}{\partial \theta^2} \cdot L + \left(\frac{\partial \ln(L)}{\partial \theta} \right)^2 \cdot L dx = 0$$

therefore

$$\left\langle - \frac{\partial^2 \ln(L)}{\partial \theta^2} \right\rangle = \left\langle \left(\frac{\partial \ln(L)}{\partial \theta} \right)^2 \right\rangle$$

Let t be an unbiased estimator of some function of θ , say $\tau(\theta)$, then

$$\begin{aligned} \langle t \rangle &= \int t L dx = \tau(\theta) \\ \tau'(\theta) &= \frac{\partial \tau(\theta)}{\partial \theta} = \int t \frac{\partial \ln(L)}{\partial \theta} L dx \end{aligned}$$

therefore from above

$$\tau'(\theta) = \int (t - \tau(\theta)) \frac{\partial \ln(L)}{\partial \theta} L dx$$

Use Schwarz inequality on $\tau'(\theta)$ to generate

$$\tau'^2 \leq \int (t - \tau)^2 L dx \times \int \left(\frac{\partial \ln(L)}{\partial \theta} \right)^2 L dx$$

Therefore, for the case $\tau(\theta) = \theta$

$$\text{var}\{t\} \geq \frac{1}{\left\langle \left(\frac{\partial \ln(L)}{\partial \theta} \right)^2 \right\rangle} = \frac{1}{\left\langle - \frac{\partial^2 \ln(L)}{\partial \theta^2} \right\rangle}$$

MALMQUIST BIAS

(also known as Eddington bias)

..... or Eddington's solution of Fredholm's integral equation of 1st kind

$$F(x) = \int U(x-z) K(z) dz$$

Expand the LH integral argument to give

$$U(x-z) = U(x) - U'(x)z + U''(x)\frac{z^2}{2!} - \dots$$

Integrate term by term

$$F(x) = \sum_n \frac{\mu_n}{n!} U^{(n)}(x)(-1)^n$$

where μ_n are moments of integration kernel K . Now rewrite

$$U(x) = F(x) + \sum_n A_n F^{(n)}(x)$$

For a central Kernel (ie. $\mu_1 = 0$) and equating coefficients

$$U(x) = F(x) - \frac{\mu_2}{2!} F^{(2)}(x) + \frac{\mu_3}{3!} F^{(3)}(x) - \left[\frac{\mu_4}{4!} - \left(\frac{\mu_2}{2!} \right)^2 \right] F^{(4)}(x) + \dots$$

For a Gaussian kernel $\mu_{odd} = 0$, $\mu_2 = \sigma^2$, $\mu_4 = 3\sigma^4$, therefore

$$U(x) = F(x) - \frac{\sigma^2}{2} F^{(2)}(x) + \frac{\sigma^4}{8} F^{(4)}(x) + \dots$$

and for say a luminosity function of the form eg. $N(m) = 10^{\alpha(m-m_o)}$

$$\frac{dN}{dm} = \ln 10 \alpha N(m) \qquad \frac{d^2N}{dm^2} = (\ln 10)^2 \alpha^2 N(m)$$

$$N_{obs}(m) = N(m) + \frac{\sigma^2}{2} (\ln 10)^2 \alpha^2 N(m) \quad \text{equivalently} \quad \Delta m = -\ln 10 \alpha \frac{\sigma^2}{2}$$